

# The effect of spatial scale on bias in regression models with spatial confounding

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# Outline

- 1 Framework
  - Correlated Residuals
  - Spatial Confounding
- 2 Results
  - Analytic
  - Simulation
- 3 Extensions

# Themes

- Intuition about residual correlation can be deceptive
- Scales of spatial correlation are critical
- Accounting for spatial correlation may help reduce bias from confounding in **some** situations.

# Uncertainty and Correlated Residuals

- Variance of regression estimates,  $\text{Var}(\hat{\beta})$ :
  - naive OLS variance is incorrect
  - GLS is the minimum variance estimator
    - lower variance than OLS with corrected variance estimate
- Question: How does residual correlation affect variance?
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# Uncertainty and Correlated Residuals

- Variance of regression estimates,  $\text{Var}(\hat{\beta})$ :
  - naive OLS variance is incorrect
  - GLS is the minimum variance estimator
    - lower variance than OLS with corrected variance estimate
- Question: How does residual correlation affect variance?
- Conventional wisdom: Correlated residuals reduce the effective sample size, so their presence adds uncertainty.

# Uncertainty and Correlated Residuals (2)

- Reality:
  - Correlated residuals offer an **opportunity** to improve precision by systematically explaining a portion of the residual variability.
  - Equivalent models

$$\text{GLS: } Y \sim \mathcal{N}(X\beta, \sigma_r^2 R + \tau^2 I)$$

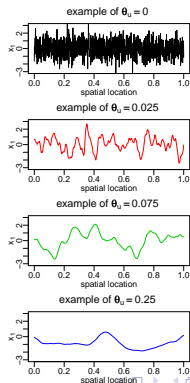
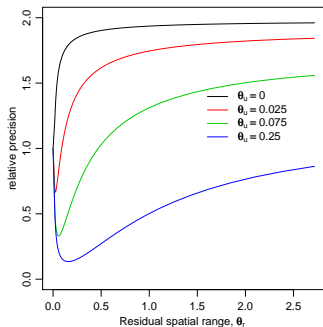
$$\begin{aligned}\text{GAM: } Y &\sim \mathcal{N}(X\beta + g, \tau^2 I) \\ g &\sim \mathcal{N}(0, \sigma_r^2 R)\end{aligned}$$

- Heuristic is that fitting either a GLS or GAM model allows one to attribute residual variability to the spatial component of the residual, reducing the unexplained variability in the model and decreasing  $\text{Var}(\hat{\beta})$ .

# Precision with Correlated Residuals

$$E(\text{Var}(\hat{\beta})^{-1}) = E(X_1^T \Sigma^{-1} X_1) = \text{tr}(\Sigma^{-1} \sigma_u^2 R(\theta_u)) = \frac{\sigma_u^2}{\tau^2} \text{tr}((I + \frac{\sigma_r^2}{\tau^2} R(\theta_r))^{-1} R(\theta_u)).$$

Results depend on the scale of the correlation in  $X_1$  and the residual.



# Key Question

- We know that attributing variability to a spatial component in the residual can reduce variance.
- Can it alleviate bias from an unmeasured, but spatially-correlated, confounder?
  - Potential mechanism: attribute variability from confounder to the spatial residual (or to a spatial term in the mean).
- Conventional Wisdom?
  - Accounting for spatial correlation in the residual can account for a spatial confounding and reduce (eliminate?) bias.
- Reality:
  - It depends on the spatial scales involved.
  - Dominici et al. (2004, JASA): control for spatial structure at large scales to eliminate confounding at that scale.
  - Goal is to assess association based on nearby observations, which share the same large-scale spatial effect.



# Thought Experiment

- Suppose pollution varies smoothly in space. Also, suppose that (unmeasured) SES varies smoothly in space.
- If we analyze a health outcome as a function of pollution, the residuals will be correlated because of SES.
- There is a fundamental non-identifiability in the model

$$Y_i = X(s_i)\beta + g(s_i) + \epsilon_i$$

which we could re-express as

$$Y_i = g^*(s_i) + \epsilon_i.$$

That is, how do we separate the pollution effect from the spatial effect (spatial confounder) if the pollution effect is just another form of spatial effect.

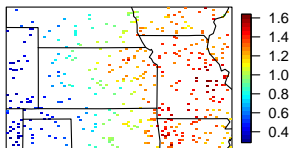
- Question: how does the model attribute variation between  $X(s)\beta$  and  $g(s)$ ?

# Scale Matters

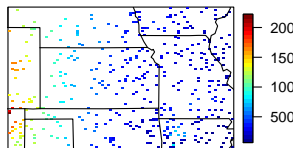
- A non-health example: how does elevation affect precipitation in the central United States?
- At large scale, precipitation increases with decreasing elevation as topography slopes gently downwards from the Rockies to the Mississippi River.
  - Elevation is not the causal effect.
- At smaller scale, precipitation increases with increasing elevation.
- A spatial model here can account for confounding from other factors that vary smoothly west to east, and isolate the elevation effect to the effect of elevation at small scales.

# Association of Elevation and Precipitation

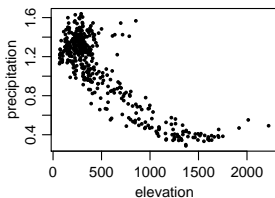
precipitation (m)



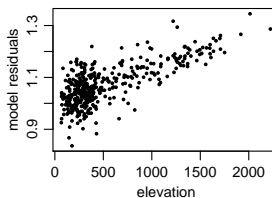
elevation (m)



Raw association



Residual association



# A Simple Model

- We can explore bias by starting with a simple generative model:

$$y_i = \beta_1 x_1(s_i) + \beta_2 x_2(s_i) + \epsilon_i$$

Let  $x_1(s)$  and  $x_2(s)$  be Gaussian processes, with  $\text{Cor}(x_1(s_i), x_2(s_i)) = \rho$ .

- If  $x_2$  is unmeasured, we arrive at the GLS model

$$y_i = \beta_1 x_1(s_i) + \epsilon_i^*$$
$$\text{Cov}(\epsilon^*) = \Sigma = \sigma_r^2 R(\theta_r) + \tau^2 I$$

where  $\sigma_r^2 = \beta_2^2 \text{Var}(x_2)$ .

- Bias comes from fitting models under the assumption that  $\epsilon^*$  is uncorrelated with  $x_1$ .

## Known parameters, single scale

- Suppose  $x_1(s)$  and  $x_2(s)$  share the same range of spatial correlation, but may be scaled differently in magnitude, namely,  $\text{Cov}(x_1) = \sigma_c^2 R(\theta_r)$  and  $\text{Cov}(\beta_2 x_2) = \beta_2^2 \sigma_2^2 R(\theta_r)$ , then

$$\begin{aligned} E(\hat{\beta}_1 | x_1) &= \beta_1 + (x_1^T \Sigma^{-1} x_1)^{-1} x_1^T \Sigma^{-1} E(x_2 | x_1) \beta_2 \\ &= \beta_1 + \rho \frac{\sigma_2}{\sigma_c} \beta_2 \end{aligned}$$

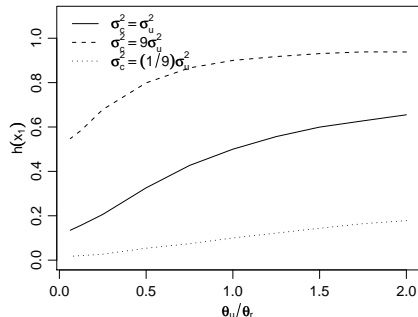
because  $E(x_2 | x_1) = \rho \sigma_2 \sigma_c R(\theta_r) \sigma_c^{-2} R(\theta_r)^{-1} x_1$ .

- The resulting bias,  $\rho \frac{\sigma_2}{\sigma_c} \beta_2$ , is the same as if the covariates were not spatially structured.
- Heuristically, the model attributes variability from the confounder to the covariate of interest.

# Known parameters, multi-scale

Let  $x_1(s) = x_c(s) + x_u(s)$  with  $\text{Cov}(x_1) = \sigma_c^2 R(\theta_r) + \sigma_u^2 R(\theta_u)$ .  
 Let  $\text{Cov}(x_2) = \sigma_c^2 R(\theta_r)$  and  $\text{Cor}(x_c(s_i), x_2(s_i)) = \rho$ .

$$\begin{aligned}
 E(\hat{\beta}_1 | x_1) & \\
 &= \beta_1 + (x_1^T \Sigma^{-1} x_1)^{-1} x_1^T \Sigma^{-1} E(x_2 | x_1) \beta_2 \\
 &= \beta_1 + \frac{x_1^T \Sigma^{-1} (I + \frac{\sigma_u^2}{\sigma_c^2} R(\theta_u) R(\theta_r)^{-1})^{-1} x_1}{(x_1^T \Sigma^{-1} x_1)} \rho \frac{\sigma_c^2}{\sigma_c} \beta_2 \\
 &= \beta_1 + h(x_1) \rho \frac{\sigma_c^2}{\sigma_c} \beta_2
 \end{aligned}$$





# Heuristics

- Reducing bias requires the covariate of interest to have a spatial scale at which it is unconfounded, and that scale must be smaller than the scale at which confounding operates.
- We would like the covariate to have as much variation at the unconfounded scale and as little at the confounded scale as possible.
- Other results are straightforward and match the non-spatial setting for confounding. We want:
  - the magnitude of variation in the confounder (or its effect on the outcome) be small.
  - the correlation between confounder and covariate to be small.



# Ongoing work

- Analysis of precision and MSE
- Simulations for non-linear settings
- Effects of choosing incorrect parameter values to minimize bias
  - Using fixed df to model the residual correlation (a la Dominici et al. 2004)
- Areal data settings
- Implications of measurement error

# Areally-aggregated Data

- Aggregated data in areal units such as zip codes, census tracts and counties are often the finest resolution data available for disease mapping analyses.
- Spatial confounding may be an issue in spatial regression models for aggregated data.
- Conditional auto-regressive (CAR) models are often used; these models smooth based on weighted averaging of neighboring units.
- Two key issues in areal models:
  - ① Aggregation smooths over fine-scale heterogeneity.
  - ② CAR models (by using local averaging) do not model large-scale spatial patterns.
- Both of these issues suggest that bias could be substantial in CAR-type models based on the results presented here.

# Measurement Error

- Classical error:
  - Preliminary work suggests that under classical error, the model attributes variability in the outcome to the spatial residual, not to the error-contaminated covariate of interest.
  - Model attenuates the effect estimate because the spatial residual is a well-measured surrogate that can stand in for the covariate.
- Berkson error/regression calibration:
  - Gryparis, Paciorek, and Coull (under revision) argue that spatial smoothing models are a form of regression calibration that induce Berkson type error when using predictions
  - Under Berkson error, we should be in the framework discussed here, except that smoothing done to make predictions will reduce fine-scale heterogeneity, decreasing our ability to reduce bias.